Behavioral/Systems/Cognitive

How Ongoing Fluctuations in Human Visual Cortex Predict Perceptual Awareness: Baseline Shift versus Decision Bias

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Visual perception fluctuates across repeated presentations of the same near-threshold stimulus. These perceptual fluctuations have often been attributed to baseline shifts—i.e., ongoing modulations of neuronal activity in visual areas—driven by top-down attention. Using magnetoencephalography, we directly tested whether ongoing attentional modulations could fully account for the perceptual impact of prestimulus activity on a subsequent seen–unseen decision. We found that prestimulus gamma-band fluctuations in lateral occipital areas (LO) predicted visual awareness, but did not reflect the focus of spatial attention. Moreover, these prestimulus signals influenced the decision outcome independently from the strength of the following visual response, suggesting that baseline shifts alone could not explain their perceptual impact. Using a straightforward decision-making model based on the accumulation of sensory evidence over time, we show that prestimulus gamma-band fluctuations in LO behave as a decision bias at stimulus onset, irrespectively of subsequent stimulus processing. In contrast, spatial attention suppressed prestimulus alpha-band signals in the same region, and produced a sustained baseline shift that also predicted the outcome of the seen–unseen decision. Together, our results indicate that prestimulus fluctuations in visual areas can influence the conscious detection of an upcoming stimulus via two distinct mechanisms: an attention-driven baseline shift in the alpha range, and a decision bias in the gamma range.

Introduction

Visual perception fluctuates across repeated presentations of the same near-threshold stimulus: sometimes the presence of a stimulus is consciously detected, and sometimes it is not. Since the two perceptual outcomes are based on the very same stimulation, these fluctuations of visual awareness are thought to reflect neuronal variability. This neuronal variability is visible in the large fluctuations of evoked sensory responses, but also in the ongoing fluctuations of prestimulus brain signals. Indeed, prestimulus fluctuations of electrophysiological (Williams et al., 2003; Linkenkaer-Hansen et al., 2004; Hanslmayr et al., 2007; Romei et al., 2008; van Dijk et al., 2008) and hemodynamic signals (Ress et al., 2000; Sapir et al., 2005; Sylvester et al., 2007; Hesselmann et al., 2008a,b) predict the conscious perception of subsequently presented stimuli. But what is the nature of these ongoing fluctuations in visual areas?

Since spatial attention can modulate visual signals in the absence of visual stimulation (Kastner et al., 1999; Siegel et al., 2008), prestimulus correlates of consciousness in visual areas have mostly been attributed to ongoing attentional modulations of sensory processing (Ress et al., 2000)—e.g., increased neuronal excitability at the focus of spatial attention (Sylvester et al., 2007). However, these so-called “baseline shifts” are not the only mechanism that can modulate ongoing brain activity and affect subsequent perceptual judgments. From a decision-making perspective, each seen–unseen report used to probe visual awareness can be considered as a perceptual decision based on the accumulation of sensory evidence about the presence of a stimulus over time (Smith and Ratcliff, 2004; Gold and Shadlen, 2007; Heekeren et al., 2008). According to this view, each perceptual decision depends on two distinct parameters: the quality of sensory evidence available in task-relevant visual areas during stimulus presentation, and the ongoing decision bias toward either percept or response at stimulus onset (Platt and Glimcher, 1999; Yang and Shadlen, 2007). Recent results suggest that some prestimulus correlates of consciousness found in visual areas could index such decision biases (Williams et al., 2003; Hesselmann et al., 2008a,b), but whether they actually reflected attention-driven baseline shifts remains largely unknown.

To address this issue, we recorded magnetoencephalographic (MEG) signals while human participants attended toward or away from low-contrast visual gratings, which were consciously detected only half of the time (Wyart and Tallon-Baudry, 2008). At the beginning of each trial (Fig. 1), participants were cued by a predictive central arrow to covertly attend either to their left or right lower visual field (Posner et al., 1980). After 600 ms, a grat-
stimuli at threshold contrast were presented either in the left or right only half of the presented stimuli were reported as present. The calibration before the experiment to estimate the threshold contrast for which background of the projection screen using a Konica Minolta LS-100 Hz) located outside the shielded recording room. The luminosity of the screen positioned at 85 cm from the participant's eyes, using a calibrated Pelli, 1997). All stimuli were presented via a mirror system on a gray tilled using the Psychtoolbox package for MATLAB (Brainard, 1997; using MATLAB (The MathWorks), and stimulus presentation was con-

fixation. The stimuli, fixation cross, and spatial cues were created off-line with their right hand.

(see Materials and Methods). Participants first selected the orientation of the grating among two choices (forced-choice discrimination task) and then reported whether they had seen a grating during the trial (detection task). We studied the perceptual impact of prestimulus MEG activity in visual areas on the subsequent seen–unseen report of visual awareness, while explicitly controlling the focus of spatial attention.

Materials and Methods

Participants

Twelve right-handed participants with normal or corrected-to-normal vision took part in the study (six women, age range 18–28 years). All participants provided informed written consent and were paid £45 for their participation. All procedures were approved by the local ethics committee (Comité Consultatif de Protection des Personnes dans la Recherche Biomédicale, Hôpital de la Pitié-Salpêtrière).

Stimuli and behavioral task

The visual stimuli were circular gratings (spatial frequency: 5.0 cycles per degree of visual angle, diameter: 2.0° of visual angle, orientation: chosen among 20 equally spaced between 0 and π, vertical and horizontal orienti-ations being excluded). The stimuli were positioned at 2.0° of visual angle from fixation with a declination of 30° in the left or right lower visual quadrants. The contrast of the stimuli was manipulated to reach threshold levels (<5%) using a simple software attenuator (Tyler, 1997). The left and right spatial cues only differed by three corners from the fixation cross and were subtending 0.9 × 0.6° of visual angle around fixation. The stimuli, fixation cross, and spatial cues were created off-line using MATLAB (The MathWorks), and stimulus presentation was controlled using the Psychotoolbox package for MATLAB (Brainard, 1997; Pelli, 1997). All stimuli were presented via a mirror system on a gray background (luminance: 26.9 cd/‌m²) at the center of a back projection screen positioned at 85 cm from the participant’s eyes, using a calibrated Mitsubishi X120 projector (resolution: 1024 × 768 pixels, refresh rate: 60 Hz) located outside the shielded recording room. The luminosity of the recording room was controlled as well as the luminance of the gray background of the projection screen using a Konica Minolta LS-100 luminance meter.

The experimental procedure has been previously described by Wyatt and Tallon-Baudry (2008). Each participant underwent a calibration session before the experiment to estimate the threshold contrast for which only half of the presented stimuli were reported as present. The calibration session was followed by eight recording sessions. Each of the eight recording sessions consisted of 92 trials, including 80 trials in which stimuli at threshold contrast were presented either in the left or right lower visual field, and 12 blank trials in which no stimulus was presented.

Of the 80 stimuli, 52 were presented at the cued location, while 28 were presented at the uncued location. Trials within a recording session were presented in a different pseudorandomized order for each participant. Participants were not given any explicit feedback regarding either the proportion of stimulus-absent trials or the presence of a stimulus at the end of each trial, since such information could have introduced undesirable prior expectations about the presence of a stimulus in subsequent trials.

MEG recordings

Continuous magnetoencephalographic signals were collected using a whole-head MEG system with 151 axial gradiometers (CTF Systems) at a sampling rate of 1250 Hz and low-pass filtered online at 300 Hz. Head localization was tracked during the experiment with respect to the MEG sensor array using marker coils that were placed at the cardinal points of the head (nasion and left and right ears). Vertical and horizontal electrooculogram (EOG) signals were simultaneously collected.

Data analysis

Data analysis was performed using in-house software (http://cogimage.dsi.cnrs. fr/logic/), the Fieldtrip package (http://www.ru.nl/fclonders/fieldtrip/), and additional programs developed in MATLAB (The MathWorks). All time samples were corrected with respect to the refresh delay of the projector (+25 ms, measured online with a photodiode).

Artifact rejection. Trials contaminated with eyes movements (rejection threshold: 1.0° of visual angle from fixation), eye blinks or muscular artifacts were rejected off-line upon visual inspection of their unfiltered EOG and MEG traces. Besides, two malfunctioning MEG sensors (MRF2 overlying right frontal cortex, and MRT3 overlying right tem- poral cortex) were discarded from all analyses.

Time–frequency analysis. A time–frequency wavelet transform was applied to each trial at each MEG sensor using a family of complex Morlet wavelets (m = 10), resulting in an estimate of oscillatory power at each time sample and at each frequency between 5 and 150 Hz. Importantly, the time–frequency resolution of the wavelets was frequency dependent (10 Hz: σ = 150 ms, 1 Hz; 50 Hz: σ = 30 ms, 5 Hz). The time–frequency data during the last 200 ms of the fixation period preceding each trial (−200 to 0 ms relative to cue onset, i.e., −800 to −600 ms relative to stimulus onset) was used as baseline, after ensuring that raw baseline power in the gamma frequency band of interest (48–64 Hz) did not differ significantly between subsequent seen and unseen reports of visual awareness (signed-rank test, p > 0.2). An index of signal power, defined at each time sample and frequency as the increase of signal power relative to fixation baseline in logarithmic units, was considered as the measure of interest for all time–frequency analyses. Indeed, the log-transformed data were distributed normally, which allowed us to use standard parametric tests (e.g., paired t test, repeated-measures ANOVA) to assess the statistical significance of observed effects.

Source estimation. To estimate the neural sources of gamma-band activity in the 50–70 Hz frequency band from a single transform, we used the “beamforming” adaptive spatial linear filtering technique (Gross et al., 2001), along with “multitaper” spectral estimates using a fixed time–frequency resolution of 30 ms, 5 Hz). The time–frequency analysis of the wavelets was frequency dependent (10 Hz: σ = 150 ms, 1 Hz; 50 Hz: σ = 30 ms, 5 Hz). The time–frequency data during the last 200 ms of the fixation period preceding each trial (−200 to 0 ms relative to cue onset, i.e., −800 to −600 ms relative to stimulus onset) was used as baseline, after ensuring that raw baseline power in the gamma frequency band of interest (48–64 Hz) did not differ significantly between subsequent seen and unseen reports of visual awareness (signed-rank test, p > 0.2). An index of signal power, defined at each time sample and frequency as the increase of signal power relative to fixation baseline in logarithmic units, was considered as the measure of interest for all time–frequency analyses. Indeed, the log-transformed data were distributed normally, which allowed us to use standard parametric tests (e.g., paired t test, repeated-measures ANOVA) to assess the statistical significance of observed effects.

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Receivers operating characteristic analysis. Receiver operating characteristic (ROC) analysis provides a measure of the degree of separation between the single-trial distributions of neuronal activities associated with seen and unseen decisions (Green and Swets, 1966). Precisely, we computed a normalized predictive power index A' derived from the area under the ROC curve. Each point of an ROC curve represents the proportion of seen decisions for which the neuronal activity was higher than a given threshold, against the proportion of unseen decisions for which

![Figure 1. Experimental design. Temporal structure of a trial, from left to right. After a variable fixation period of 0.8–1 s, participants were cued by a predictive central arrow to covertly attend either to their left or right lower visual field. After 600 ms, a low-contrast grating at detection threshold appeared for 400 ms either at the cued or uncued location, or occasionally did not appear at all. Participants first selected the orientation of the grating among two choices and then reported whether they believed a grating was presented during the trial, by pressing one of two buttons with their right hand.](https://example.com/figure1.png)
the neuronal activity exceeded the same threshold. The ROC curve can be traced by varying the threshold between the minimum and maximum values of both distributions. We used a nonparametric estimate of the area under the ROC curve corresponding to the trapezoidal area under the empirical ROC curve (Macmillan and Creelman, 2005). The predictive power $A'$ corresponded to a scaled percentile version of the estimate of the area under the ROC curve (AUC) between $-1$ and $+1$:

$$A' = 2 \cdot (\text{AUC} - 0.5).$$

This index can be interpreted as the objective performance of an ideal observer at predicting the outcome of the seen–unseen decision based only on neuronal activity for a single trial. A null predictive power indicates that the corresponding neuronal activity does not predict the seen–unseen decision better than chance, whereas a positive or negative value indicates that the corresponding neuronal activity predicts the subsequent seen–unseen decision better than chance, with higher or lower levels of neuronal activity for predicted seen decisions, respectively.

**Extraction of residual variability.** We used partial rank correlations between single-trial distributions of prestimulus and poststimulus frequency-specific activities at the participant level to extract the following: (1) the residual variability in the prestimulus time window ($-300$ to $-100$ ms relative to stimulus onset) which did not correlate with trial-to-trial fluctuations in the poststimulus time window ($+250$ to $+450$ ms relative to stimulus onset), and (2) the residual variability in the poststimulus time window which did not correlate with trial-to-trial fluctuations in the prestimulus time window. We verified that there was no spurious correlation between the residual prestimulus variability and the fluctuations in the poststimulus time window for the gamma (rank correlation test, mean $p = +0.05, p > 0.2$ for all participants) and alpha (rank correlation test, mean $p = +0.04, p > 0.2$ for all participants) bands.

**Decision-making model.** We tested the model predictions regarding the predictive power of gamma- and alpha-band activity on the seen–unseen decision using computational simulations ($n = 10,000$ simulations). For each model simulation, we computed for each trial of each participant the model decision based on the observed prestimulus and poststimulus frequency-specific activities. We then sorted the single-trial distributions of prestimulus and poststimulus activities relative to the model-based seen and unseen decisions, and computed the corresponding predictive power of the prestimulus and poststimulus activities on model-based decisions. We used these two values (averaged across computational simulations) to obtain unique estimates for the two free parameters in the model: $w_{\gamma}$ corresponding to the relative strength of the decision bias in the seen–unseen decision (in $[0, 1]$ range), and $w_{\text{sec}}$ corresponding to the absolute influence of gamma-band activity in the decision (also in $[0, 1]$ range). Once the estimates were obtained, we computed the model prediction regarding the predictive power of the residual prestimulus variability, and compared it to the observed value for participants. We defined the model error as the discrepancy between the mean model prediction and the observed value (in normal deviate Z-score units).

**Results**

Perceptual fluctuations did not only reflect fluctuations of attention

Behavioral results have been previously described by Wyart and Tallon-Baudry (2008). We report here the results relevant for the present study. At detection threshold (contrast $= 3.3 \pm 0.3\%$, mean $\pm$ SEM), participants reported the presence of a stimulus in half of stimulus presentations (hit rate $= 46.6 \pm 2.5\%$) but much less often in blank trials (false-alarm rate $= 14.9 \pm 2.6\%$), corresponding to a detection $d'$ of $1.04 \pm 0.14$ ($t$ test against zero, $p < 0.001$) along with a conservative decision criterion $\beta$ of $+0.69 \pm 0.14$ ($t$ test against zero, $p < 0.001$) (Green and Swets, 1966). Participants correctly discriminated the orientation of the grating in $67.2 \pm 1.9\%$ of stimulus presentations, corresponding to a discrimination $d'$ of $0.90 \pm 0.10$ ($t$ test against zero, $p < 0.001$). Moreover, participants discriminated the orientation of the presented grating much more accurately when they consciously detected its presence (detected gratings: $84.5 \pm 3.2\%$, undetected gratings: $51.7 \pm 0.9\%$, signed-rank test, $p < 0.001$).

As expected (Posner et al., 1980; Pashler et al., 2001), spatial attention improved the speed and accuracy of orientation discrimination judgments (signed-rank test, mean reaction time to detected gratings: attentional cueing effect $= -47 \pm 13$ ms, $p < 0.01$, correct discrimination rate: attentional cueing effect $= +4.9 \pm 2.6\%$, $p < 0.05$). Precisely, mean reaction times were $1101 \pm 62$ ms to cued detected gratings, $1148 \pm 67$ ms to cued undetected gratings, $1110 \pm 91$ ms to cued undetected gratings, and $1121 \pm 98$ ms to misjudged undetected gratings. However, spatial attention had only a moderate effect on visual awareness, since participants detected $49.9 \pm 3.1\%$ of cued gratings, against as much as $40.4 \pm 3.0\%$ of misjudged gratings (signed-rank test, $p < 0.05$). Participants put as much effort into detecting and discriminating the orientation of the gratings: detection $d'$ and discrimination $d''$ did not differ significantly (paired $t$ test, $p > 0.1$), and the attentional cueing effects on the detection task (hit rate: cued minus misjudged) and the discrimination task (correct discrimination rate: cued minus misjudged) were positively correlated across participants (rank correlation test, $p = +0.77, n = 12, p < 0.005$), suggesting similar attentional constraints for the two tasks.

The sequence of perceptual reports of visual awareness appeared random. First, there were no more repetitions of seen and unseen decisions than predicted by chance ($x^2$ goodness-of-fit test against binomial distribution, $p > 0.3$). Second, participants were not biased in their decision criterion $\beta$ by their previous seen–unseen choice (paired $t$ test, $p > 0.5$), or by the presence (or absence) of a stimulus in the previous trial (paired $t$ test, $p > 0.5$). Finally, detection sensitivity remained constant over the four time quartiles of the experiment (repeated-measures ANOVA, hit rate: $p > 0.1$, detection $d'$: $p > 0.1$).

Altogether, the results suggest that fluctuations of perceptual reports reflected neither prior expectations about the presence of a stimulus based on the previous trial nor slow fluctuations of alertness. Besides, the focus of spatial attention could account only partially for these trial-to-trial fluctuations. Therefore, there remained a large part of unexplained variability in the perceptual reports of visual awareness.

From neural correlates to neural predictors of visual awareness

The analysis of the oscillatory MEG response to the near-threshold visual stimuli revealed a neural dissociation between correlates of visual awareness and spatial attention in human visual cortex (Wyart and Tallon-Baudry, 2008). In particular, gamma-band activity increased in response to detected stimuli but was not modulated by spatial attention. Interestingly, this activity correlated with seen–unseen decisions even in blank trials—i.e., in the absence of visual stimulation. We therefore suspected that ongoing fluctuations of gamma-band activity before stimulus onset could also influence the conscious detection of the near-threshold stimuli. To test this hypothesis, we extended the analysis of the oscillatory MEG activity to the prestimulus time period, at the time when participants focused their attention at the cued location, over the same occipital region of interest (ROI) that correlated with visual awareness during stimulus presentation (Fig. 2A). During the prestimulus time period, occipital MEG signals showed frequency-specific modulations relative to fixation (Fries et al., 2008; Siegel et al., 2008); power suppressions at low frequencies, in particular in the alpha range, and power increases at high frequencies in the gamma range (Fig. 2B).

Prestimulus fluctuations of gamma-band activity ($48–64$ Hz)
over the occipital ROI correlated with the conscious detection of upcoming stimuli at \(-300\) to \(-100\) ms relative to stimulus onset (repeated-measures ANOVA, \(p < 0.01\)) (Fig. 2C). Although prestimulus gamma-band signals were equally strong over left and right occipital ROI across trials (\(p > 0.5\)), their perceptual impact was spatially selective, in the sense that only prestimulus fluctuations contralateral to the location of the upcoming stimulus differed depending on the subsequent seen–unseen decision (double interaction: \(p < 0.05\), contralateral to the upcoming stimulus; \(p > 0.3\)) (Fig. 3A). This spatial selectivity suggests that the prestimulus gamma-band activity was localized in retinotopic visual areas.

But did these prestimulus fluctuations of gamma-band activity reflect the focus of spatial attention? Surprisingly, spatial attention did not modulate prestimulus gamma-band activity contralaterally to the cued location (repeated-measures ANOVA, \(p > 0.5\)) (Fig. 3B). Besides, the direction of the cue did not influence the effect of prestimulus gamma-band activity on the subsequent seen–unseen decision (double interaction: \(p > 0.2\)). To further exclude the possibility that this spatially selective prestimulus effect on visual awareness is driven by the direction of the cue, we assessed that the perceptual impact of prestimulus gamma-band activity on visual awareness was more dependent on the location of the upcoming stimulus than on the direction of the cue (repeated-measures ANOVA, triple interaction: \(p < 0.05\)).

Nevertheless, spatial attention induced frequency-specific modulations during the prestimulus time period, but at lower frequencies. In accordance with previous results (Thut et al., 2006; Fries et al., 2008; Siegel et al., 2008), the alpha-band suppression (10–12 Hz) over the occipital ROI was stronger contralaterally to the cued location at \(-300\) to \(-100\) ms relative to stimulus onset (repeated-measures ANOVA, \(p < 0.001\)). The temporal dynamics of this attentional modulation of alpha-band activity differed markedly from the dynamics of the prestimulus effect of gamma-band activity on visual awareness (Fig. 2D).

Therefore, prestimulus gamma-band fluctuations in visual areas predicted subsequent perceptual reports of visual awareness in a spatially selective manner, but they did not reflect the focus of spatial attention, in contrast to prestimulus alpha-band fluctuations.

Source estimates of gamma-band activity in lateral occipital areas

We confirmed these findings in source space. We first estimated the sources of gamma-band activity (50–70 Hz) during the presentation of a detected stimulus (+250 to +450 ms relative to stimulus onset), using spatial filtering techniques (see Materials and Methods). The source estimates of gamma-band activity were localized in extrastriate visual areas overlapping the lateral occipital gyrus, which we will refer to as lateral occipital areas (LO). We defined a bilateral ROI at the peak of gamma-band activity (>50% of the maximum relative power across all voxels, Talairach coordinates: \(x = -37/+33, y = -84, z = -16\) mm) (Fig. 4A, left). This high-frequency response in LO was stronger contralaterally to detected stimuli (repeated-measures ANOVA, \(p < 0.05\)) (Fig. 4A, right), correlated with visual awareness (\(p < 0.005\), but was not modulated by the focus of spatial attention (\(p > 0.5\)), as previously observed in sensor space (Wyart and Tallon-Baudry, 2008).

Prestimulus gamma-band fluctuations in LO differed depending on the subsequent seen–unseen decision at \(-300\) to \(-100\) ms relative to stimulus onset (repeated-measures ANOVA, \(p < 0.05\)) (Fig. 4B). As observed in sensor space, the difference was significant only for the hemisphere contralateral to the upcoming stimulus (double interaction: \(p < 0.05\), contralateral to the upcoming stimulus: \(p > 0.005\), ipsilateral to the upcoming stimulus: \(p > 0.5\)). During the same prestimulus time period, spatial attention induced frequency-specific modulations in LO in the alpha range (repeated-measures ANOVA, \(p < 0.001\)), but not in the gamma range (repeated-measures ANOVA, \(p > 0.1\)).

These results show that the prestimulus gamma-band fluctuations that predicted the subsequent seen–unseen decision were localized in the same visual region (LO) and in the same frequency range as the following visual response to the near-
To measure quantitatively the perceptual impact of gamma-band activity, we computed the predictive power of gamma-band activity on the detected/unnoticed decision. But how did the prestimulus and poststimulus gamma-band activity differ depending on the subsequent report of visual awareness (**p < 0.01; ns, nonsignificant). Gamma-band activity is expressed in relative power. Error bars indicate SEM.

**Figure 3.** Spatial selectivity of prestimulus gamma-band activity. A. Spatial selectivity of prestimulus gamma-band effect on visual awareness. Prestimulus gamma-band activity (48–64 Hz, −300 to −100 ms relative to stimulus onset) for detected and undetected stimuli, over the occipital ROI contralateral and ipsilateral to the location of the upcoming stimulus. Only contralateral prestimulus fluctuations differed depending on the subsequent report of visual awareness (**p < 0.01; ns, nonsignificant). Gamma-band activity is expressed in relative power. Error bars indicate SEM. B. Absence of attentional modulation of prestimulus gamma-band activity. Prestimulus gamma-band activity for attended and unattended hemispheres, over the occipital ROI contralateral and ipsilateral to the location of the upcoming stimulus. The focus of spatial attention did not modulate the pattern of prestimulus gamma-band activity.

**Figure 4.** Source estimates of gamma-band activity. A, ROI in source space. Left. The source estimates of the visual gamma-band response to the stimuli (50–70 Hz, 250 to 450 ms relative to stimulus onset) were found in LO (Talairach coordinates in millimeters; LH, left hemisphere; RH, right hemisphere). White lines correspond to the displayed slices. Right, The visual gamma-band response peaked contralaterally to detected stimuli (left panel, contralateral hemisphere; right panel, ipsilateral hemisphere). Black arrows indicate the contralateral ROI. Gamma-band activity is expressed in relative power. B. Time course of gamma-band activity in LO. Prestimulus gamma-band activity in LO differed depending on the subsequent report of visual awareness. Time is expressed relative to stimulus onset. Gamma-band activity is expressed in relative power. Error bars indicate SEM. Thick lines at the bottom indicate when gamma-band activity differs significantly depending on detected (aware) and undetected (unaware) stimuli (black, p < 0.01; gray, p < 0.05).

Threshold stimulus—which also correlated with the seen/unseen decision. But how did the prestimulus and poststimulus gamma-band fluctuations contribute to perceptual awareness?

**Perceptual impact of prestimulus gamma-band activity**

To measure quantitatively the perceptual impact of gamma-band activity, we computed the predictive power $A'$ of gamma-band activity on the seen/unseen decision at a single-trial level using ROC analysis (Green and Swets, 1966, see Materials and Methods). This bias-free index can be interpreted as the performance of an ideal observer at predicting the participant’s seen/unseen decision based only on gamma-band activity at a single time point of a single trial. Positive and negative values correspond to higher and lower levels of gamma-band activity for predicted seen decisions, respectively. In sensor space, the predictive power of gamma-band activity peaked in two distinct time windows over the occipital ROI contralateral to the stimulus (Fig. 5A): a prestimulus time window at $t = −300$ to −100 ms relative to stimulus onset ($t$ test against zero, $A' = +6.2 ± 1.5, p < 0.001$) and a poststimulus time window at 250–450 ms after stimulus onset ($t$ test against zero, $A' = +7.5 ± 1.3, p < 0.001$).

We applied the same approach to blank trials—i.e., trials for which no stimulus was presented. In these trials, participants sometimes erroneously detected the presence of a physically absent stimulus. Although there were only a few of these false-alarm reports per participant (from $n = 1$ to $n = 26$, median $n = 7$), the predictive power of gamma-band activity over the bilateral occipital ROI for blank trials revealed a qualitatively similar pattern with two successive peaks, one before and one after stimulus onset. The first predictive peak, centered around −100 ms relative to stimulus onset, almost reached statistical significance ($t$ test against zero, $p = 0.06$), suggesting that prestimulus gamma-band fluctuations in visual areas influenced subsequent seen/unseen decisions even when no stimulus was presented.

In stimulus trials, the predictive power of gamma-band activity was not different before and after stimulus onset (paired $t$ test, $p > 0.2$), but showed a significant drop in a peristimulus time window centered at stimulus onset ($±100$ ms) (repeated-measures ANOVA, $p = 0.001$). Indeed, the predictive power in the peristimulus time window was not different from zero ($t$ test against zero, $A' = +0.8 ± 1.8, p > 0.5$) and was significantly weaker than in the prestimulus (post hoc comparison, $p < 0.005$) and poststimulus time windows (post hoc comparison, $p < 0.005$). Besides, this significant drop in predictive power around stimulus onset did not merely reflect an overall decrease of gamma-band activity between the prestimulus and peristimulus time windows (paired $t$ test, $p > 0.5$).

To test whether prestimulus and poststimulus gamma-band fluctuations reflected the same variability, we computed the correlation between the trial-to-trial fluctuations of gamma-band activity in the prestimulus and poststimulus time windows. The correlation between these two time windows was significant (rank correlation test, mean $p = +0.49, p < 0.001$ for all participants). This observation points toward the existence of a baseline shift—i.e., an ongoing level of gamma-band activity which affected both prestimulus and poststimulus time windows. This
positive correlation could indicate that the perceptual impact of prestimulus gamma-band activity on the seen–unseen decision relies only on the strength of the following gamma-band response. In other words, prestimulus gamma-band fluctuations could predict visual awareness only because they reflected ongoing fluctuations that also modulated the strength of the subsequent gamma-band response to the near-threshold stimulus—i.e., only because of the presence of a baseline shift.

To test for this interpretation, we extracted from prestimulus gamma-band fluctuations the residual variability which did not correlate with the fluctuations in the poststimulus time window (and vice versa for the residual poststimulus variability). The residual prestimulus gamma-band variability still predicted the seen–unseen decision better than chance (**p < 0.001, *p < 0.01, *p < 0.05). Error bars indicate SEM.

Interpretation within the perceptual decision-making framework

Altogether, our results suggest that prestimulus fluctuations of gamma-band activity in visual areas biased directly perceptual decisions, independently from the strength of the follow-

ing gamma-band response. Because these prestimulus fluctuations precede the gamma-band response by several hundreds of milliseconds, they must leave a neuronal “trace” at the decision level. This idea is well accounted for in the perceptual decision-making framework (Smith and Ratcliff, 2004; Gold and Shadlen, 2007; Heekeren et al., 2008). According to this view, each seen–unseen decision is based on the accumulation of sensory evidence about the presence of a stimulus over time. Qualitatively, our results fit well the perceptual decision-making framework and suggest that (1) gamma-band activity is integrated at the decision level as sensory evidence during stimulus presentation (Siegel et al., 2007); and (2) this accumulation of sensory evidence is biased at stimulus onset by prestimulus gamma-band fluctuations.

To formally test this interpretation of our findings, we designed a straightforward decision-making model of the seen–unseen decision based on gamma-band fluctuations in visual areas (Fig. 6A). Precisely, we computed a synthetic decision variable corresponding to a noisy observation of the weighted sum of prestimulus and poststimulus gamma-band activities. At the single-trial level, the weighted poststimulus gamma-band activity is considered as sensory evidence about the presence of a stimulus, while the weighted prestimulus gamma-band activity is considered as an ongoing decision bias toward the seen (or unseen) decision at stimulus onset. Formally, the decision variable $X_{\text{dec}}$ is described by the following:

$$X_{\text{dec}} = w_{\text{dec}} \cdot Z(w_0 \cdot X_{\text{pre}} + (1 - w_0) \cdot X_{\text{post}}) + (1 - w_{\text{dec}}) \cdot N_{(0,1)},$$

where $X_{\text{pre}}$ is the vector of prestimulus gamma-band activities across trials, $X_{\text{post}}$ is the vector of poststimulus gamma-band activities across trials, $Z(\cdot)$ is the normal deviate, and $N_{(0,1)}$ is Gaussian noise with zero mean and unit variance. There are two free parameters in the model: $w_0$ corresponds to the relative strength of the decision bias in the seen–unseen decision (in $[0, 1]$ range), while $w_{\text{dec}}$ corresponds to the absolute influence of gamma-band activity in the decision (also in $[0, 1]$ range). For each trial of each participant, we used the estimates of $X_{\text{pre}}$ and $X_{\text{post}}$ from the experimental data to compute $X_{\text{dec}}$. Following signal detection theory (Green and Swets, 1966), if $Z(x_{\text{dec}})$ exceeded $\Phi^{-1}(h)$, where $\Phi^{-1}(\cdot)$ is the inverse cumulative normal distribution function and $h$ is the hit rate of the participant across trials, then the seen decision was taken. Otherwise, the unseen decision was taken.

We used the single-trial distributions of prestimulus and poststimulus gamma-band activities to compute model-based perceptual decisions at the single-trial level (see Materials and Methods). Therefore, the observed baseline shift—i.e., the correlation between prestimulus and poststimulus gamma-band activities—was also present in model simulations. If $w_0 = 0$, then the prestimulus gamma-band activity does not influence directly the decision-making process—i.e., there is no decision bias. In contrast, if $w_0 > 0$, then the prestimulus gamma-band activity influences directly the decision-making process by producing a decision bias.

We first sampled the whole parameter space by computing for each possible $(w_{\text{dec}}, w_0)$ couple the predictive power of prestimulus and poststimulus gamma-band activities on model-based perceptual decisions using ROC analysis (Fig. 6B). The observed mean predictive power of gamma-band activity for participants uniquely defined estimates for $w_0$ and $w_{\text{dec}}$ ($w_0 = 0.35$, $w_{\text{dec}} = 0.08$). The estimate for $w_0$ was noticeably larger than zero.

Having determined the values of the model parameters $w_0$ and $w_{\text{dec}}$, we computed the decision variable $X_{\text{dec}}$ for each participant and each trial, and then we compared the model-based perceptual decisions with the actual decisions using ROC analysis (Fig. 6B). The observed mean predictive power of gamma-band activities for participants uniquely defined estimates for $w_0$ and $w_{\text{dec}}$ ($w_0 = 0.35$, $w_{\text{dec}} = 0.08$). The estimate for $w_0$ was noticeably larger than zero.

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Figure 6. Interpretation within the perceptual decision-making framework. A, Decision-making model. The prestimulus gamma-band activity (PRE box) is followed by the poststimulus gamma-band activity (POST box). In turn, the poststimulus gamma-band activity is accumulated as sensory evidence about the presence of a stimulus during the decision process (DEC box). The prestimulus gamma-band activity influences subsequent perceptual decisions either indirectly via its correlation with the poststimulus gamma-band activity (black arrow), or directly as a decision bias (blue arrow). The extraction of the residual prestimulus variability (light gray bar) from the total prestimulus variability (dark gray bar) allows to choose between these two possible interpretations. B, Estimation of model parameters. We computed the predictive power $A'$ of gamma-band activity on model-based decisions for each possible ($w_o$, $w_{dec}$) couple in the prestimulus (left panel) and poststimulus (right panel) time windows. Isolines correspond to the observed predictive power of gamma-band activity on participants’ decisions in the prestimulus ($A' = 6.2$) and poststimulus ($A' = 7.5$) time windows. Both isolines are shown on each panel. Their intersection uniquely defined estimates for the model parameters $w_o$ and $w_{dec}$. $w_o = 0.35$, $w_{dec} = 0.08$. C, Fitted model. Top, Schematic view of the decision-making model. The prestimulus gamma-band activity influences directly the subsequent model-based decision as a decision bias ($w_{dec} = 0.35$). Strong gamma-band increases lead to seen decisions. Bottom, Model predictions regarding the predictive power $A'$. Thick horizontal lines correspond to the fitted values from the observed values for participants used to obtain estimates for $w_o$ and $w_{dec}$. Center error bars indicate SEM. Side error bars indicate the discrepancy between model predictions and the observed values for participants (model error $> 1$). Model predictions did not match the observed values (ns, nonsignificant).

$w_{dec}$ based on the predictive power of the prestimulus and poststimulus gamma-band activities, we then tested the model prediction regarding the predictive power of the residual prestimulus gamma-band variability on the seen–unseen decision (see Materials and Methods). The model prediction was strikingly close to the observed value, both qualitatively and quantitatively (Fig. 6C). Indeed, the predicted $A'$ was significantly positive (one-tailed $t$ test, median $p < 0.05$) and was equal to $+3.1$, compared with $+3.1 \pm 1.4$ for participants (model error $< 0.1$, one-tailed $t$ test against model prediction, $p > 0.4$). More generally, the model predictions revealed the same qualitative effects observed for participants (Table 1).

To determine whether the quality of the model fit critically depended on the presence of a decision bias by prestimulus gamma-band fluctuations, we suppressed it in subsequent model simulations by setting $w_o$ to zero (while keeping the previous estimate for $w_{dec} = 0.08$). This time, the model predictions did not match observed values for participants, especially regarding the predictive power of the residual prestimulus gamma-band activity on perceptual decisions (Fig. 6D). The predicted $A'$ was no longer different from zero (one-tailed $t$ test, median $p > 0.4$) and was equal to $+0.4$ compared with $+3.1 \pm 1.4$ for participants (model error $= 1.6$, one-tailed $t$ test against model prediction, $p < 0.05$) (Table 1).

In both simulations, the amount of correlation between prestimulus and poststimulus gamma-band activities was the same as in the participants’ data. The only difference between the two models is the presence (or absence) of a decision bias on the subsequent seen–unseen decision. We compared the two models using the Bayes factor (Kass and Raftery, 1995)—i.e., the ratio of the likelihood of observing the participants’ values given each of the two models ($w_o = 0.35$ and $w_o = 0$, respectively). The model with the decision bias was 18 times more likely ($+2.9$ in natural logarithmic units) than the model without the decision bias.

Therefore, a straightforward decision-making model of the seen–unseen decision based on the integration of gamma-band fluctuations in visual areas could successfully account for participants’ decisions. Importantly, the model predictions were found to be accurate only if prestimulus gamma-band fluctuations produced a decision bias on the subsequent seen–unseen decision, independently from the strength of the following gamma-band response.
Perceptual impact of prestimulus alpha-band activity

Prestimulus gamma-band fluctuations in visual areas did not reflect the focus of spatial attention. Indeed, spatial attention did not modulate the observed pattern of prestimulus gamma-band activity and did not interfere with the effect of prestimulus gamma-band activity on the seen–unseen decision. Nevertheless, spatial attention also influenced visual awareness, since participants detected the presence of a stimulus more often when presented at the cued location. To determine whether this attentional cueing effect also corresponded to a decision bias at stimulus onset, we used the same model-based approach as above to study how prestimulus correlates of spatial attention influenced subsequent seen–unseen decisions.

Since spatial attention induced a frequency-specific suppression of prestimulus alpha-band signals in visual areas (Thut et al., 2006; Fries et al., 2008; Siegel et al., 2008; Wyart and Tallon-Baudry, 2008), we computed the predictive power \( A' \) of alpha-band activity on the seen–unseen decision over time (Fig. 7A). Interestingly, alpha-band activity started to predict the subsequent seen–unseen decision better than chance in the prestimulus time window, when spatial attention started to suppress alpha-band activity contralaterally to the cued location, over the occipital ROI contralateral to the upcoming stimulus (\( t \) test against zero, \( A' = -3.3 \pm 1.5, p < 0.05 \)). The predictive power gradually increased and peaked in the poststimulus time window (\( t \) test against zero, \( A' = -8.0 \pm 1.6, p < 0.001 \)). In contrast to the gamma range, the predictive power of alpha-band activity on perceptual decisions was consistently stronger after than before stimulus onset (paired \( t \) test, \( p < 0.001 \)). Nevertheless, the correlation between prestimulus and poststimulus alpha-band fluctuations (rank correlation test, mean \( \rho = +0.49, p < 0.001 \)) was as strong as the one previously observed within the gamma range (paired \( t \) test, \( p > 0.5 \)), indicating the presence of a baseline shift in the alpha range. However, there was no correlation between alpha- and gamma-band fluctuations in either the prestimulus (rank correlation test, mean \( \rho = -0.01, p > 0.05 \)) or post-stimulus (rank correlation test, mean \( \rho = -0.01, p > 0.05 \)) time window—i.e., mutually independent baseline shifts in the alpha and gamma ranges.

As before, we extracted from prestimulus alpha-band fluctuations the residual variability which did not correlate with post-stimulus fluctuations in the same frequency range, and computed its predictive power on the seen–unseen decision using ROC analysis. In contrast to what was observed in the gamma range, the residual prestimulus alpha-band variability no longer predicted the seen–unseen decision better than chance (\( t \) test against zero, \( A' = +0.4 \pm 0.8, p > 0.5 \)) (Fig. 7B). This result suggests that prestimulus fluctuations of alpha-band activity did not produce a decision bias on the seen–unseen decision.

We confirmed this interpretation using the same decision-making model as before (Fig. 6A). This time, we used the single-trial distributions of prestimulus and poststimulus alpha-band suppressions to compute model-based perceptual decisions at the single-trial level, and obtained unique estimates for \( w_0 \) and \( w_{\text{dec}} \) (\( w_0 = 0.00, w_{\text{dec}} = 0.08 \)). Importantly, the estimate for \( w_0 \) was zero, and the estimate for \( w_{\text{dec}} \) was the same as in the gamma range. We tested the model prediction regarding the predictive power of the residual prestimulus alpha-band variability on the seen–unseen decision. The model prediction matched the observed value for participants (Fig. 7C). The predicted \( A' \) was not different from zero (one-tailed \( t \) test, median \( p > 0.4 \)) and was equal to \(-0.4\), compared with \(+0.4 \pm 1.0 \) for participants (model error \( =0.5 \), one-tailed \( t \) test against model prediction, \( p > 0.2 \)) (Table 1). To determine whether a model with a decision bias could better fit participants’ decisions, we added such a bias in subsequent model simulations by setting \( w_0 \) to 0.35 (the estimate for \( w_0 \) in the gamma range). The corresponding model predictions did not match the observed values for participants, especially regarding the predictive power of the residual prestimulus alpha-band variability (model error \( =2.1 \), one-tailed \( t \) test against model prediction, \( p < 0.005 \)) (Table 1). Again, we compared the two models using the Bayes factor and found that the model without a decision bias was 58 times more likely (+4.1 in logarithmic units) than the model with a decision bias as strong as the one observed in the gamma range.

In conclusion, the observed prestimulus correlates of spatial attention in visual areas did not produce any decision bias on the subsequent seen–unseen decision. In contrast to what was observed in the gamma range, their perceptual impact could be explained solely by the presence of a sensory baseline shift—i.e., ongoing modulations that also affected the strength of the following alpha-band response to the near-threshold stimulus.

**Discussion**

We described here prestimulus alpha- and gamma-band fluctuations in visual areas which both predict the conscious fate of
subsequently presented visual stimuli, but yet correspond to two distinct mechanisms. Prestimulus correlates of consciousness in visual areas have mostly been related to attention-driven baseline shifts (Ress et al., 2000; Sylvester et al., 2007). We confirmed this interpretation by reporting attentional suppressions of prestimulus alpha-band signals in contralateral lateral occipital areas (Thut et al., 2006; Fries et al., 2008; Siegel et al., 2008; Wyart and Tallon-Baudry, 2008). This temporal discontinuity of the predictive power suggests that prestimulus and poststimulus gamma-band fluctuations in LO reflect qualitatively distinct sensory processes. While poststimulus gamma-band fluctuations are thought to reflect fluctuations of stimulus processing per se (Siegel et al., 2007), prestimulus gamma-band fluctuations might correspond to spontaneous sensory predictions about the upcoming stimulus (Hessemann et al., 2008a,b)—e.g., the preactivation of a sensory “template”—independently of spatial attention.

The perceptual impact of gamma-band activity in LO on the seen–unseen decision fits a general perceptual decision-making framework (Smith and Ratcliff, 2004; Gold and Shadlen, 2007; Heekeren et al., 2008) in which prestimulus and poststimulus gamma-band fluctuations in LO could reflect a decision bias on the accumulation of sensory evidence over time, irrespectively of subsequent stimulus processing.

Our results highlight the distinction between two types of prestimulus effects on subsequent perceptual decisions. We show that sensory baseline shifts and decision biases can co-occur independently in the same region, in distinct frequency bands. The predictive power of the alpha-band suppression observed in LO built up progressively from cue onset throughout the trial (Romei et al., 2008), suggesting a sustained modulation of neuronal excitability by spatial attention which affected the processing of the near-threshold stimulus. In contrast, the predictive power of gamma-band activity in LO showed two successive peaks, one before and one after stimulus onset, with independent contributions to subsequent perceptual decisions.

The two predictive peaks are separated by a ~250 ms time period, centered around stimulus onset, during which gamma-band activity in LO does not correlate with the subsequent seen–unseen decision. This observation underlines that gamma-band signals in LO do not reflect the decision variable per se, whose predictive power on the following perceptual choice should increase monotonically over time (Williams et al., 2003). In theory, this nonpredictive time period could be reminiscent of the “nondecision time”—i.e., the time before the accumulation process begins (Ratcliff, 1978). The existence of a nondecision time for stimulus encoding is postulated in most decision-making models to account for the distribution of response times in simple perceptual tasks (Smith and Ratcliff, 2004). Interestingly, the nonpredictive time period observed here corresponds approximately to the delay before gamma-band activity starts to reflect stimulus-specific processing in contralateral LO (Wyart and Tallon-Baudry, 2008).

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tion process into a single decision variable. However, our results only suggest (rather than demonstrate) the existence of an accumulation of sensory evidence over time. We do not know whether this accumulation process maps onto a single brain area, such as the intraparietal cortex (Williams et al., 2003; Tosoni et al., 2008), or onto a mosaic of processes and regions (Heekeren et al., 2004; Philia-stides and Sajda, 2007). Furthermore, the decision bias observed here in LO appears distinct from previously described decision priors based on stimulus or reward expectations (Platt and Glimcher, 1999; Yang and Shadlen, 2007). Indeed, such probabilistic priors have been found at the decision level in frontoparietal regions. Here, in contrast, the observed decision bias arises from ongoing fluctuations at the sensory level in retinotopic visual areas.

The two types of prestimulus effects reported here—the attention-driven baseline shift in the alpha range and the decision bias in the gamma range—both predict perceptual awareness, but as distinct contributions to the same decision process. Curiously, the search for the neural correlates of consciousness only occasionally considers explicitly the perceptual decision-making framework as a useful model of how subjective judgments might be implemented in the brain (Resset al., 2000; Supèr et al., 2001; Williams et al., 2003; Hesselmann et al., 2008a,b). However, our results highlight the qualitative distinction between attentional and decision mechanisms of consciousness. Adopting this distinction not only accounts for the present results, but also sheds light on previous findings. First, the seen–unseen report used to probe visual awareness in laboratory experiments is known to be intrinsically influenced by other cognitive functions, such as attention and working memory (Lamme, 2006). This complexity could potentially be teased apart by considering subjective reports as the outcome of a single decision process which integrates independent sensori-motor functions, which might be differentially influenced by awareness, attention and working memory. Second, neural correlates of visual consciousness have previously been found both in occipital (Supèr et al., 2001; Rees and Heeger, 2003; Wilke et al., 2006; Wyatt and Tallon-Baudry, 2008) and frontoparietal regions (Dehaene et al., 2001; Shulman et al., 2001; Sergent et al., 2005). The present results suggest that the relevant perceptual information available in visual areas is gradually integrated over time into a decision variable that might map onto frontoparietal regions (de Lafuente and Romo, 2006; Goldberg et al., 2006). From this decision-making perspective, occipital and frontoparietal correlates of visual consciousness thus correspond to distinct parts of the same decision process, with different temporal dynamics and different top-down influences from other cognitive functions (Gilbert and Sigman, 2007).

Finally, our results suggest that visual awareness critically depends on the accumulation of sensory information in visual areas over time. In that sense, the corresponding decision variable therefore constitutes a neuronal “trace” of the recent sensory past, interpreted according to ongoing expectations and task goals. Such a mnemonic trace could potentially explain the perceived continuity of visual consciousness—i.e., its link with a form of brief sensory memory (Atkinson et al., 2000; Baars and Franklin, 2003). More generally, the accumulation of noisy sensory information in visual areas over time provides a generic mechanism that could account for the perceptual stability of conscious visual experience.

References


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